Lecture 2: Uniform Convergence and Optimization

Ruhr-University Bochum

April 29, 2019

▲□▶ ▲□▶ ▲三▶ ▲三▶ 三三 のへで

Concentration inequalities: Nuts and bolts

• Markov:

$$\mathbb{P}\Big(\big|X\big| \ge \varepsilon\Big) \le \frac{\mathbb{E}\big|X\big|}{\varepsilon}$$

・ロト・西ト・ヨト・ヨー うへぐ

Concentration inequalities: Nuts and bolts

• Markov:

$$\mathbb{P}\Big(|X| \ge \varepsilon\Big) \le \frac{\mathbb{E}|X|}{\varepsilon}$$

• Chebychev:

$$\mathbb{P}\Big(\frac{|X - \mathbb{E}X|}{\sqrt{\operatorname{var}(X)}} \ge \varepsilon\Big) \le \frac{1}{\varepsilon^2}$$

◆□▶ ◆□▶ ◆臣▶ ◆臣▶ 善臣 - のへで

Concentration inequalities: Nuts and bolts

Markov:

$$\mathbb{P}\Big(\big|X\big| \ge \varepsilon\Big) \le \frac{\mathbb{E}\big|X\big|}{\varepsilon}$$

• Chebychev:

$$\mathbb{P}\Big(\frac{|X - \mathbb{E}X|}{\sqrt{\operatorname{var}(X)}} \ge \varepsilon\Big) \le \frac{1}{\varepsilon^2}$$

• Chernov: X_i iid, in [0, 1]

$$\mathbb{P}\Big(\sum_{i=1}^n X_i - \mathbb{E}X_i \ge \varepsilon\Big) \le \exp(-2\varepsilon^2/(2n))$$

▲□▶ ▲□▶ ▲三▶ ▲三▶ 三三 のへで

- A statistical query is a (measurable) function $\phi : \mathcal{X} \to [0, 1]$.
- Notation: \mathcal{D} is some probability distribution on \mathcal{X} . Population mean:

$$\phi(\mathcal{D}) := \mathbb{E}_{x \sim \mathcal{D}} \phi(x)$$

 $S \sim D^n$ iid data sample consisting of $X_1, ..., X_n$. Empirical mean:

$$\phi(\mathcal{S}) = \frac{1}{n} \sum_{i=1}^{n} \phi(X_i)$$

・ロト ・ 日 ・ モート ・ モー・ つくぐ

How do we model the situation of interest?

- A's aim: Find out features of \mathcal{D} : $\phi_1(\mathcal{D}), ..., \phi_K(\mathcal{D})$.
- Problem: Neither \mathcal{D} is known, nor is \mathcal{S} directly accessible.
- A interacts with a mechanism M, which returns answers $a_1, ..., a_K$ to his queries.

How do we model the situation of interest?

- A's aim: Find out features of \mathcal{D} : $\phi_1(\mathcal{D}), ..., \phi_K(\mathcal{D})$.
- \bullet Problem: Neither ${\cal D}$ is known, nor is ${\cal S}$ directly accessible.
- A interacts with a mechanism M, which returns answers $a_1, ..., a_K$ to his queries.

If M is in charge of some data base with sensitive information, A will not always get all the information.

How do we model the situation of interest?

- A's aim: Find out features of \mathcal{D} : $\phi_1(\mathcal{D}), ..., \phi_K(\mathcal{D})$.
- \bullet Problem: Neither ${\cal D}$ is known, nor is ${\cal S}$ directly accessible.
- A interacts with a mechanism M, which returns answers $a_1, ..., a_K$ to his queries.

If M is in charge of some data base with sensitive information, A will not always get all the information. Examples of M:

- 1. Empirical mechanism: Returns $\phi_j(S)$ for ϕ_j .
- 2. Privatizing mechanism: Returns $\phi_j(S) + L_j$ for ϕ_j . L_j is some Laplace noise.

A theoretical measure of performance is given by

$$\textit{err}_{\mathcal{S}}(M,A) = \max_{j=1,...,K} |\phi_j(\mathcal{D}) - a_j|$$

We say M is (α, β) -accurate for K queries on iid data for every analyst A if

$$\mathbb{P}(\operatorname{err}_{\mathcal{S}}(M,A) \leq \alpha) \geq 1 - \beta.$$

Sometimes we look at expected errors such as

$$\sup_{\mathcal{D}} \sup_{A} \mathbb{E}err_{\mathcal{S}}(M, A).$$

or even

$$\inf_{M} \sup_{\mathcal{D}} \sup_{A} \mathbb{E}err_{\mathcal{S}}(M, A).$$

The next Theorem demonstrates that a high accuracy is feasable via the empirical mechanism M_{emp} .

Theorem 3:

Let \mathcal{D} be any probability distribution on $\mathcal{X}, \phi_1, ..., \phi_K$ (data independent) statistical queries of the analyst A and $S \sim \mathcal{D}^n$ iid data. Then with probability $\geq 1 - \delta$

$$\operatorname{err}_{\mathcal{S}}(M_{\operatorname{emp}},A) \leq \sqrt{rac{\log(2K/\delta)}{2n}}$$

イロト 不得 トイヨト イヨト ヨー ろくで

Proof:

Last time we have seen that with probability $\geq 1-\alpha$ for any query

$$|\phi(\mathcal{S}) - \phi(\mathcal{D})| \leq \sqrt{rac{\log(2/lpha)}{2n}}.$$

Choosing $\alpha = \delta/K$ we see that

$$\mathbb{P}\Big(\exists j \in \{1,...,K\} : |\phi_j(\mathcal{S}) - \phi_j(\mathcal{D})| > \sqrt{rac{\log(2K/\delta)}{2n}}\Big) \ \leq \sum_{j=1}^{K} \mathbb{P}\Big(|\phi_j(\mathcal{S}) - \phi_j(\mathcal{D})| > \sqrt{rac{\log(2K/\delta)}{2n}}\Big) \leq \delta.$$

◆□▶ ◆□▶ ◆臣▶ ◆臣▶ 善臣 - のへで

Corollary:

Under the Assumptions of Theorem 3:

$$\mathbb{E}err_{\mathcal{S}}(M_{emp}, A) = \mathcal{O}\left(\sqrt{\frac{\log(2K)}{2n}}\right)$$

◆□▶ ◆□▶ ◆ □▶ ◆ □▶ ● □ ● ● ● ●

Statistical models are often determined by a parameter $w \in \Theta \subset \mathbb{R}^d$, which can be expressed as the minimizer of an averaged loss function, i.e.

$$w^* = \operatorname{argmin}_{w \in \Theta} \mathbb{E}_{x \sim \mathcal{D}} \ell(w, x).$$

How do we find out w^* ? A typical estimator is the empirical minimizer

$$\hat{w}_{emp} = argmin_{w \in \Theta} \sum_{i=1}^{n} \ell(w, X_i).$$

This srategy can only be successful if ℓ is "well behaved" in some fashion.

Assumptions:

- $\Theta \subset \{u : \|u\| \leq R\}.$
- For all $x \in \mathcal{X}$

$$|\ell(u;x) - \ell(v;x)| \leq |u-v|C \quad u,v \in \Theta.$$

Theorem 5:

Under the above assumption for iid data $\mathcal{S} \sim \mathcal{D}^n$ it holds with probability $\geq 1 - \delta$:

$$\sup_{u\in\Theta} |\sum_{i=1}^n \ell(u;X_i) - \mathbb{E}_{x\sim\mathcal{D}}\ell(u;x)| \le 6RC\sqrt{\frac{d\log(n/\delta)}{n}}.$$

Theorem 5 suggests minimizing the empirical loss

$$\hat{w}^* = \operatorname{argmin}_w \sum_{i=1}^n \ell(w; X_i).$$

How difficult is minimizing this?

For the broad class of convex optimization problems this task is efficiently solvable.

Subdifferential:

The subdifferential of a function $f : \mathbb{R}^d \supset \Theta \to \mathbb{R}$ in a point x is defined as

$$\partial f(x) := \{g \in \mathbb{R}^d : f(x) + \langle g, y - x \rangle \leq f(y) \forall y \in \Theta \}.$$

Convex optimization

Convexity:

- A set Θ ⊂ ℝ^d is called convex if it equals a (possibly infinite) intersection of halfspaces.
- 2 Let Θ ⊂ ℝ^d be a convex set. A function f : Θ → ℝ is convex on Θ iff ∂f(x) ≠ Ø for all x ∈ Θ.

Remarks:

Let Θ be closed and convex. The projection

$$\Pi_{\Theta}(x) := \operatorname{argmin}_{w \in \Theta} \|x - w\|$$

is well defined for all $x \in \mathbb{R}^d$. Furthermore for all $w \in \Theta$ and $y \in \mathbb{R}^d$ projection reduces distances, i.e.

$$\|\Pi_{\Theta}(y) - w\| \le \|y - w\|$$

Projected Gradient Descent

Let $f : \Theta \to \mathbb{R}$ be a function and Θ a convex subset of \mathbb{R}^d . The method of PGD is defined as follows:

1. Choose some $x_0 \in \Theta$, $\eta > 0$ and $T \in \mathbb{N}$.

2. Set
$$y_{t+1} = x_t - \eta g_t$$
, where $g_t \in \partial f(x_t)$

3. Set
$$x_{t+1} = \Pi_{\Theta}(y_{t+1})$$
.

 If t + 1 = T stop and output x_T. Else set t = t + 1 and repeat 2.

Theorem 10:

Let $\Theta \subset \{u : ||u|| \le R\}$ be closed, convex and $f : \Theta \to \mathbb{R}$ be convex and *C*-lipschitz. If we run PGD *T* times with $\eta = R/(C\sqrt{T})$, then

$$f\left(\frac{1}{T}\sum_{t=1}^{T}x_{t}\right)-f(x^{*})\leq RC/\sqrt{T},$$

where $x^* \in \Theta$ is the minimizer of f.